

5.26

$$x(t) = \sum_{n=-\infty}^{\infty} (-1)^n p(t-n)$$

$$X(t) = x(t-D), \quad D \sim \text{uniform } [0, 1].$$

(a)

$$R_{xx}(t_1, t_2) = E\{x(t_1)x(t_2)\}$$

$$= E\{x(t_1-D)x(t_2-D)\}$$

$$= E\left\{\sum_{n=-\infty}^{\infty} (-1)^n p(t_1-n-D) \cdot \sum_{m=-\infty}^{\infty} (-1)^m p(t_2-m-D)\right\}$$

$$= \sum_{n=-\infty}^{\infty} \sum_{m=-\infty}^{\infty} (-1)^{n+m} E\{p(t_1-n-D)p(t_2-m-D)\}$$

$$E\{p(t_1-n-D)p(t_2-m-D)\} = \int_0^1 p(t_1-n-a)p(t_2-m-a) da$$

$$\text{let } t_1-n-a = u$$

$$\Rightarrow t_2-m-a$$

$$\Rightarrow -da = du$$

$$= t_2-m-(t_1-n-u)$$

$$= (t_2-t_1)-(m-n)+u$$

$$= - \int_{t_1-n}^{t_1-n-1} p(u) p((t_2-t_1)-(m-n)+u) du$$

$$= \int_{t_1-n-1}^{t_1-n} p(u) p(u + \underbrace{(t_2-t_1)}_{\tau} - (m-n)) du$$

$$\therefore R_{xx}(t_1, t_2) = \sum_{n=-\infty}^{\infty} \sum_{m=-\infty}^{\infty} (-1)^{n+m} \int_{t_1-n-1}^{t_1-n} p(u) p(u + \tau - (m-n)) du$$

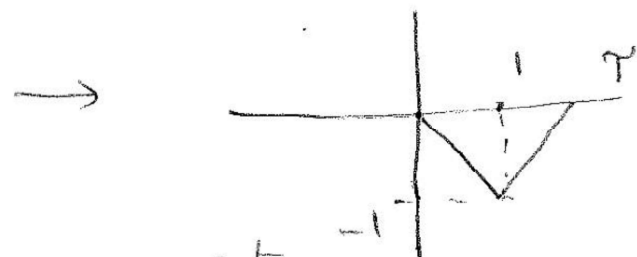
for $m=n$

$$R_{xx}(t_1, t_2) = \sum_{m=-\infty}^{\infty} \int_{t_1-m-1}^{t_1-m} p(u) p(u+\tau) du$$

$$= \int_{-\infty}^{\infty} p(u) p(u+\tau) du \rightarrow$$

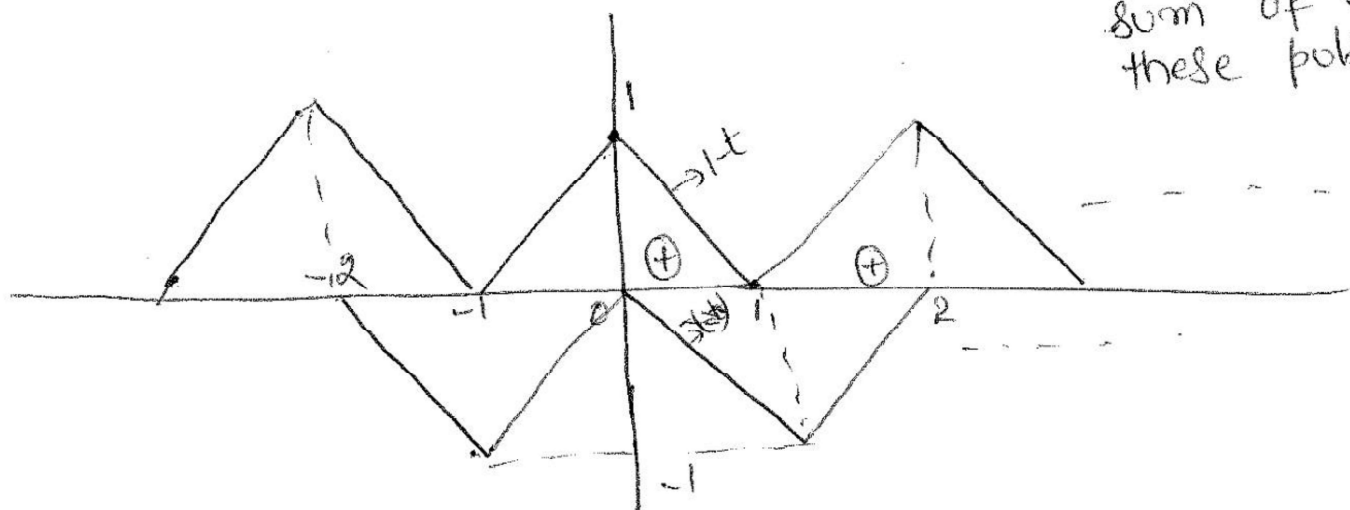
for $(m-n)=1$

$$R_{xx}(t_1, t_2) = \sum_{m=-\infty}^{\infty} (-1)^m \int_{t_1-(m+1)-1}^{t_1-(m+1)} p(u) p(u+\tau-1) du$$



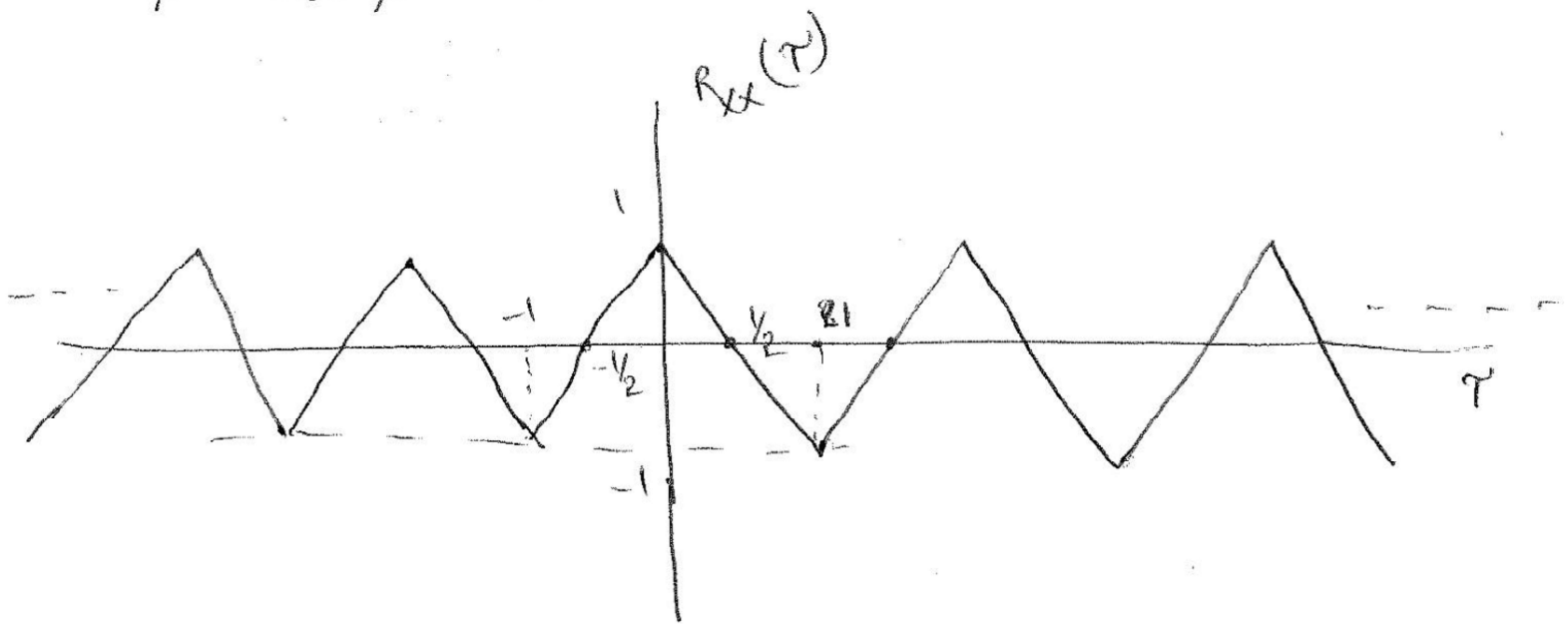
proceeding in a similar way we get

$$R_{xx}(t_1, t_2) = R_{xx}(\tau)$$



sum of all these pulses

This results in



(b).

$$E\{x(t)\} = E\{x(t-0)\}$$

$$= E\left\{\sum_{n=-\infty}^{\infty} (-1)^n p(t-n-0)\right\}$$

$$= \sum_{n=-\infty}^{\infty} (-1)^n E\{p(t-n-0)\}$$

$$= \sum_{n=-\infty}^{\infty} (-1)^n \cdot \int_0^1 p(t-n-a) \cdot 1 \cdot da$$

$$= \sum_{n=-\infty}^{\infty} (-1)^n \cdot \int_0^1 I_{[-1/2, 1/2]}(t-n-a) da.$$

$$\text{let } t-n-a = u$$

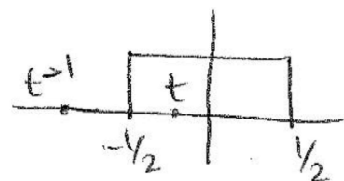
$$= \sum_{n=-\infty}^{\infty} (-1)^n \int_{t-n}^{t-n-1} I_{[-1/2, 1/2]}(u) (-du)$$

$$= \sum_{n=-\infty}^{\infty} (-1)^n \int_{t-(n+1)}^{t-n} I_{[-1/2, 1/2]}(u) du$$

for $n=0$

$$= \int_{t-1}^t I_{[-1/2, 1/2]}(u) du$$

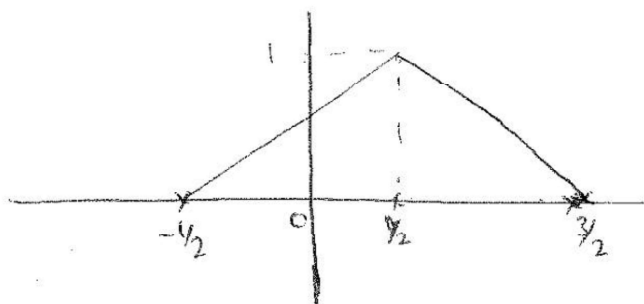
$$= \int_{-1/2}^t du = t + 1/2 \quad \text{if } -1/2 < t < 1/2$$



$$= \int_{t-1}^{\frac{1}{2}} du \quad \text{if} \quad \frac{1}{2} < t < \frac{3}{2}$$

$$= \frac{1}{2} - (t-1)$$

$$= \frac{3}{2} - t$$



f(2) $n=1$

$$\int_{-1/2}^{t-1} dt$$

$$-1/2 < t-1 < 1/2$$

$$1/2 < t < 3/2$$

$$= t-1 + 1/2$$

$$= t - 1/2$$

$$= \int_{t-2}^{\frac{1}{2}} dt \quad \text{if}$$

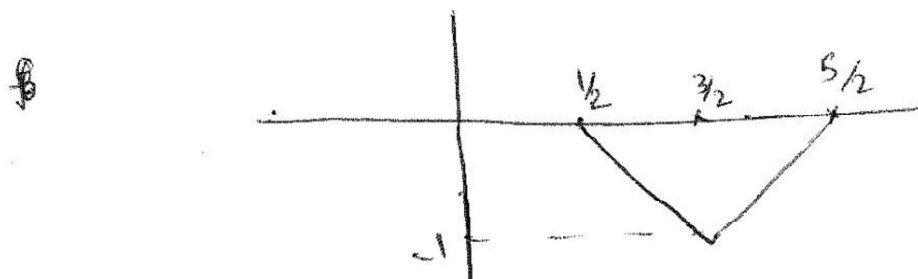
$$-1/2 < t-2 < 1/2$$

$$3/2 < t < 5/2$$

$$= \frac{1}{2} - (t-2)$$

$$= -t + 3/2$$

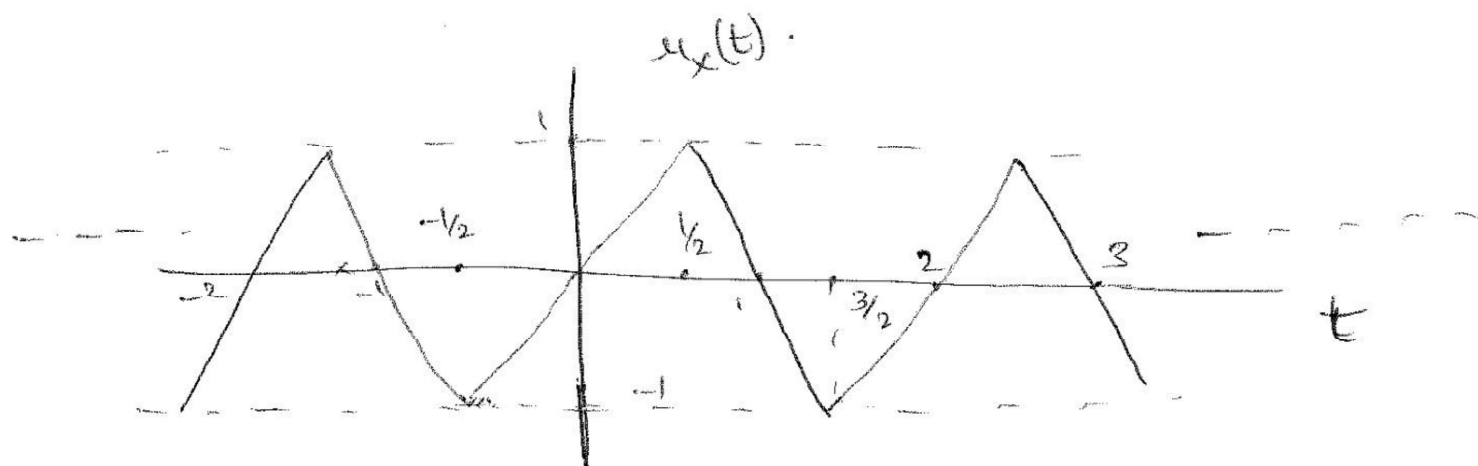
(simply shifted by 1)



(sign reversed because of $(-y)^n$)

iii) for other n

by adding all these we get

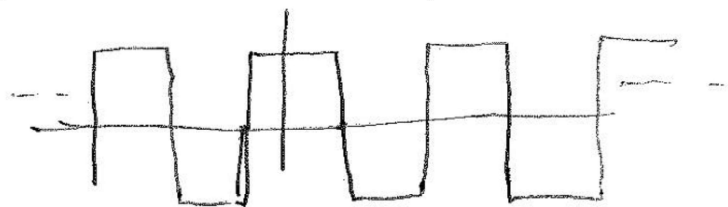


(b) $\therefore \mu_x(t)$ depends on t it is not WSS

(c) $\therefore x$ is not WSS it is not stationary

(d) consider time average of mean

$$\bar{\mu}_x(t) = \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T x(t-d) dt$$



over infinite intervals mean = 0

$$\bar{\mu}_x(t) = 0$$

$\therefore x$ is not ergodic in mean.

from soln to 5.205 and 5.206(a)

we see that ensemble autocorrelation is same as time averaged autocorrelation.

So x is ergodic in autocorrelation.

Pdf approach to find mean

$x(t)$ at $t=t_1$ can take either $+1$ or -1 .

$$P\{x(t_1) = 1\}$$

$$= P\{x(t_1 - D) = 1\}$$

$$= P\{-\frac{1}{2} \leq t_1 - D \leq \frac{1}{2}\}$$

$$= P\{-\frac{1}{2} \leq D - t_1 \leq \frac{1}{2}\}$$

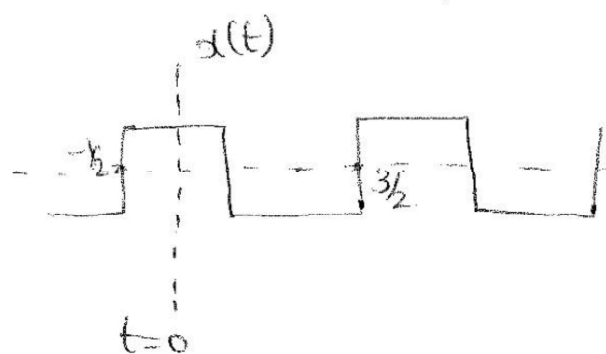
$$= P\{t_1 - \frac{1}{2} \leq D \leq t_1 + \frac{1}{2}\}$$

□

$$\text{if } 0 \leq t_1 + \frac{1}{2} \leq 1$$

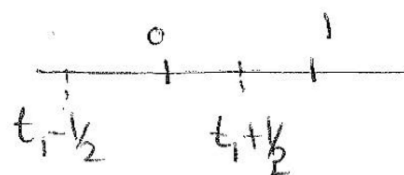
$$-\frac{1}{2} \leq t_1 \leq \frac{1}{2}$$

$$\begin{aligned} \text{then } P\{t_1 - \frac{1}{2} \leq D \leq t_1 + \frac{1}{2}\} &= (t_1 + \frac{1}{2} - 0) \cdot 1 \\ &= t_1 + \frac{1}{2} \end{aligned}$$



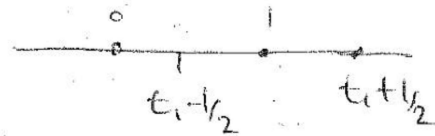
consider $-\frac{1}{2} \leq t_1 \leq \frac{3}{2}$

$x(t)$ is periodic with period 2.



if $0 \leq t_1 - \frac{1}{2} \leq 1$

$\Rightarrow \frac{1}{2} \leq t_1 \leq \frac{3}{2}$



$$P\left\{t_1 - \frac{1}{2} \leq D \leq t_1 + \frac{1}{2}\right\} = \left[1 - \left(t_1 - \frac{1}{2}\right)\right] \cdot 1$$

$$= \frac{3}{2} - t_1$$

$$\therefore P\{x(t_1)=1\} = \begin{cases} t_1 + \frac{1}{2} & -\frac{1}{2} \leq t_1 \leq \frac{1}{2} \\ \frac{3}{2} - t_1 & \frac{1}{2} \leq t_1 \leq \frac{3}{2} \end{cases}$$

$$E\{x(t_1)\} = 1 \cdot P\{x(t_1)=1\} - 1 \cdot P\{x(t_1)=-1\}$$

$$= P\{x(t_1)=1\} - \{1 - P\{x(t_1)=1\}\}$$

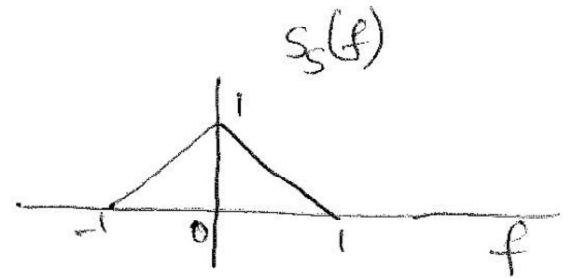
$$= 2P\{x(t_1)=1\} - 1$$

$$\therefore E\{x(t_1)\} = \begin{cases} 2t_1 - 1 & -\frac{1}{2} \leq t_1 \leq \frac{1}{2} \\ 3 - 2t_1 & \frac{1}{2} \leq t_1 \leq \frac{3}{2} \end{cases}$$

$\therefore x(t)$ is periodic this results in same answer as previous.

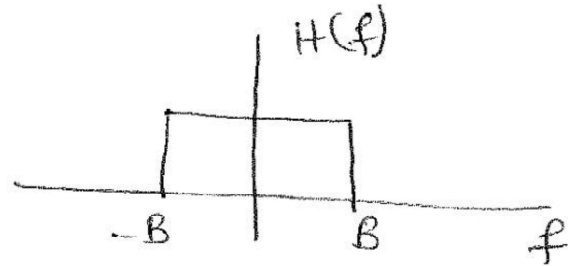
5.31

$$S_s(f) = (1 - |f|) \mathbb{I}_{[-1, 1]}(f)$$



$$y(t) = s(t) + n(t)$$

$$S_n(f) \equiv 0.001$$



(a) SNR at filter i/p

$$= \frac{\text{signal power}}{\text{noise power}}$$

$$\text{signal power} = \int_{-1}^1 S_s(f) df = \frac{1}{2} \times 2 \times 1 = 1$$

$$\text{noise power} = 2 \times 0.001 = 0.002$$

$$\text{SNR} = \frac{1}{2 \times 10^{-3}} = 500$$

(b) for $B=1$

$$\text{SNR}_{\text{o/p}} = \text{SNR}_{\text{i/p}}$$

for $B = \frac{1}{2}$

$$\begin{aligned} \text{Signal power at the o/p} &= 2 \int_0^{\frac{1}{2}} (1-f) df \\ &= \frac{3}{4} \end{aligned}$$

$$\text{noise power} = 10^{-3}$$

$$\therefore \text{SNR}_{o/p} = \frac{3}{4} \times 1000$$
$$= 750$$

$\therefore$ SNR is better for $B = 1/2$.

Problem 6.1

(a) Let $Z = (y * h)(t_0)$, where $h(t) = s(-t)$. Then $Z \sim N(m, v^2)$ if 1 sent, and $Z \sim N(0, v^2)$ if 0 sent, where

$$v^2 = \sigma^2 \|h\|^2 = \sigma^2 4 \int_0^1 t^2 dt = \frac{4}{3} \sigma^2$$

$$m = (s * h)(t_0) = \int s(t)s(t - t_0)dt = \int_0^1 t(1 - t)dt = \frac{1}{6}$$

Thus, for the ML decision rule

$$P_e = Q\left(\frac{|m|}{2v}\right) = Q\left(\frac{1}{8}\sqrt{\frac{E_b}{N_0}}\right)$$

using the fact that we must have

$$\frac{|m|}{2v} = a\sqrt{\frac{E_b}{N_0}}$$

(why?) where a is a constant determined by substituting $E_b = \frac{1}{2}\|s\|^2 = \frac{2}{3}$ and $N_0 = 2\sigma^2$.

(b) We can improve the error probability by sampling at $t_0 = 0$. We then have $m = \|s\|^2 = \frac{4}{3}$, while v^2 is as before. This gives, reasoning as in (a),

$$P_e = Q\left(\frac{|m|}{2v}\right) = Q\left(\sqrt{\frac{E_b}{N_0}}\right)$$

the usual formula for the performance of optimal reception of on-off keying in AWGN.

(c) For $h(t) = I_{[0,2]}$, we again have the same model for the decision statistic $Z = y * h(t_0)$, but with $v^2 = \sigma^2 \|h\|^2 = 2\sigma^2$, and $m = (s * h)(t_0)$. The performance improves with $|m|$, which is maximized at $t_0 = 2$ ($m = 1$) or $t_0 = 4$ ($m = -1$). We therefore get that, for the ML decision rule,

$$P_e = Q\left(\frac{|m|}{2v}\right) = Q\left(\sqrt{\frac{3E_b}{8N_0}}\right)$$

(d) Note that we can approximate the matched filter $s(-t)$ using linear combinations of two shifted versions of $h(t) = I_{[0,2]}$, by approximating triangles by rectangles. That is, the matched filter shape is approximated as $\tilde{h}(t) = h(t+2) - h(t+4)$. Thus, we can use the decision statistic

$$Z = (y * \tilde{h})(0) = (y * h)(2) - (y * h)(4)$$

We now have $Z \sim N(m, v^2)$ if 1 sent, and $Z \sim N(0, v^2)$ if 0 sent, where

$$v^2 = \sigma^2 \|\tilde{h}\|^2 = 4\sigma^2$$

$$m = (s * \tilde{h})(0) = 2$$

As before, we can enforce the scaling with $\frac{E_b}{N_0}$ to get

$$P_e = Q\left(\frac{|m|}{2v}\right) = Q\left(\sqrt{\frac{3E_b}{4N_0}}\right)$$